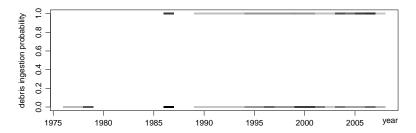
# M-estimation when data values are not completely known

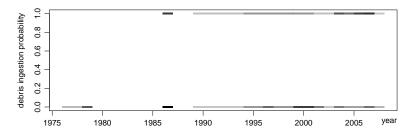
Marco Cattaneo

Department of Mathematics University of Hull

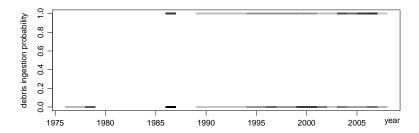
CFE-CMStatistics 2015, London, UK 13 December 2015



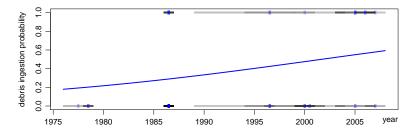
 for 468 green turtles, the data describe the presence or absence of marine debris in the gastrointestinal system at the time of death, which is interval-censored (Schuyler et al., 2014)



- for 468 green turtles, the data describe the presence or absence of marine debris in the gastrointestinal system at the time of death, which is interval-censored (Schuyler et al., 2014)
- the main question is if the probability of debris ingestion increased over time: we decide to use logistic regression to answer it

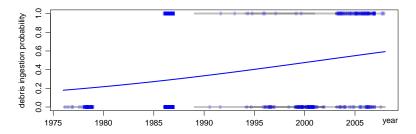


how can we deal with incomplete/censored data?



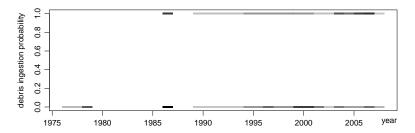
how can we deal with incomplete/censored data?

delete them or make them precise (e.g., interval midpoints) and apply the conventional statistical method: easy, but what does the result mean?



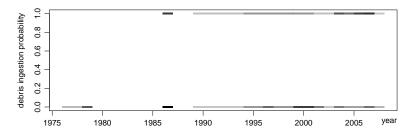
how can we deal with incomplete/censored data?

- delete them or make them precise (e.g., interval midpoints) and apply the conventional statistical method: easy, but what does the result mean?
- impute the precise values and apply the conventional statistical method: the result depends on the imputation assumptions (e.g., uniform on intervals)



how can we deal with incomplete/censored data?

- delete them or make them precise (e.g., interval midpoints) and apply the conventional statistical method: easy, but what does the result mean?
- impute the precise values and apply the conventional statistical method: the result depends on the imputation assumptions (e.g., uniform on intervals)
- apply the conventional statistical method to all compatible precise data sets: why? and how?



how can we deal with incomplete/censored data?

- delete them or make them precise (e.g., interval midpoints) and apply the conventional statistical method: easy, but what does the result mean?
- impute the precise values and apply the conventional statistical method: the result depends on the imputation assumptions (e.g., uniform on intervals)
- apply the conventional statistical method to all compatible precise data sets: why? and how?
- adapt the conventional statistical method to the case of incomplete/censored data: we will follow this approach

• data:  $X_1, \ldots, X_n \in \mathcal{X}$  i.i.d.

- data:  $X_1, \ldots, X_n \in \mathcal{X}$  i.i.d.
- M-estimator minimising the nonparametric MLE of  $E[\rho(X_i, \theta)]$ :

$$\hat{\theta}(X_1,\ldots,X_n) = \arg\min_{\theta} \sum_{i=1}^n \rho(X_i,\theta)$$

- data:  $X_1, \ldots, X_n \in \mathcal{X}$  i.i.d.
- M-estimator minimising the nonparametric MLE of  $E[\rho(X_i, \theta)]$ :

$$\hat{\theta}(X_1,\ldots,X_n) = \arg\min_{\theta} \sum_{i=1}^n \rho(X_i,\theta)$$

 nonparametric/pragmatic approach: under weak regularity conditions (Huber and Ronchetti, 2009),

$$\hat{\theta}(X_1,\ldots,X_n) \xrightarrow[n \to \infty]{a.s.} \arg\min_{\theta} E[\rho(X_i,\theta)]$$

- data:  $X_1, \ldots, X_n \in \mathcal{X}$  i.i.d.
- M-estimator minimising the nonparametric MLE of  $E[\rho(X_i, \theta)]$ :

$$\hat{\theta}(X_1,\ldots,X_n) = \arg\min_{\theta} \sum_{i=1}^n \rho(X_i,\theta)$$

 nonparametric/pragmatic approach: under weak regularity conditions (Huber and Ronchetti, 2009),

$$\hat{\theta}(X_1,\ldots,X_n) \xrightarrow[n \to \infty]{a.s.} \arg\min_{\theta} E[\rho(X_i,\theta)]$$

▶ parametric/idealistic approach: assuming further  $X_i \sim P_\theta$  and  $E[\rho(X_i, \theta)] < E[\rho(X_i, \theta')]$ ,

$$\hat{\theta}(X_1,\ldots,X_n) \xrightarrow[n \to \infty]{a.s.} \theta$$

▶ data:  $S_1, \ldots, S_n \subseteq \mathcal{X}$  i.i.d., with  $X_i \in S_i$  unknown

- ▶ data:  $S_1, \ldots, S_n \subseteq \mathcal{X}$  i.i.d., with  $X_i \in S_i$  unknown
- M-estimator minimising the nonparametric MLE of  $E[\rho(X_i, \theta)]$ :

$$\hat{\theta}(S_1, \dots, S_n) \stackrel{?}{=} \arg\min_{\theta} \operatorname{co}\left\{\sum_{i=1}^n \rho(x_i, \theta) \, : \, x_i \in S_i\right\}$$

- ▶ data:  $S_1, \ldots, S_n \subseteq \mathcal{X}$  i.i.d., with  $X_i \in S_i$  unknown
- M-estimator minimising the nonparametric MLE of E[ρ(X<sub>i</sub>, θ)]:

$$\hat{ heta}(S_1,\ldots,S_n) \stackrel{?}{=} \arg\min_{\theta} \operatorname{co}\left\{\sum_{i=1}^n \rho(x_i,\theta) \, : \, x_i \in S_i\right\}$$

nonparametric/pragmatic approach: under weak regularity conditions,

$$\hat{\theta}_{\min\max}(S_1,\ldots,S_n) \xrightarrow[n \to \infty]{a.s.} \arg\min_{\theta} \overline{E}[\rho(X_i,\theta)],$$

where  $\overline{E}[\rho(X_i, \theta)]$  is the maximum/supremum expectation compatible with the distribution of  $S_i$ 

- ▶ data:  $S_1, \ldots, S_n \subseteq \mathcal{X}$  i.i.d., with  $X_i \in S_i$  unknown
- M-estimator minimising the nonparametric MLE of E[ρ(X<sub>i</sub>, θ)]:

$$\hat{\theta}(S_1,\ldots,S_n) \stackrel{?}{=} \arg\min_{\theta} \operatorname{co}\left\{\sum_{i=1}^n \rho(x_i,\theta) \, : \, x_i \in S_i\right\}$$

nonparametric/pragmatic approach: under weak regularity conditions,

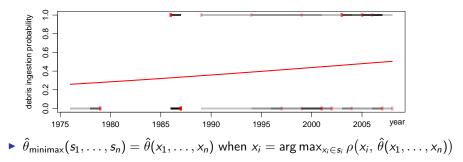
$$\hat{\theta}_{\min\max}(S_1,\ldots,S_n) \xrightarrow[n \to \infty]{a.s.} \arg\min_{\theta} \overline{E}[\rho(X_i,\theta)],$$

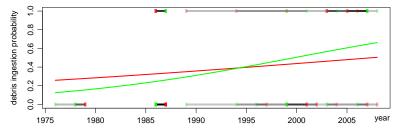
where  $\overline{E}[\rho(X_i, \theta)]$  is the maximum/supremum expectation compatible with the distribution of  $S_i$ 

parametric/idealistic approach: under additional assumptions,

$$\hat{\theta}_{\min\min}(S_1,\ldots,S_n) \xrightarrow[n \to \infty]{a.s.} [\theta],$$

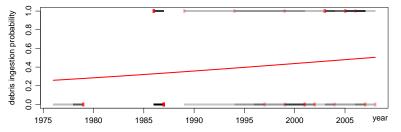
where  $[\theta]$  is the identification region of  $\theta$  (Manski, 2003): smoothing corrections (of the  $\varepsilon$ -minimin form) may be needed in case of partial identification





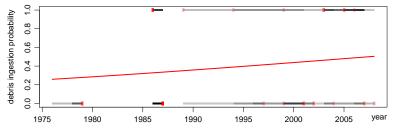
 $\bullet \ \hat{\theta}_{\min\max}(s_1,\ldots,s_n) = \hat{\theta}(x_1,\ldots,x_n) \text{ when } x_i = \arg\max_{x_i \in s_i} \rho(x_i, \ \hat{\theta}(x_1,\ldots,x_n))$ 

the minimax logistic regression with one interval-censored covariate can always be obtained by computing at most two conventional logistic regressions (for the two extreme choices of compatible precise data sets)



 $\bullet \ \hat{\theta}_{\min\max}(s_1,\ldots,s_n) = \hat{\theta}(x_1,\ldots,x_n) \text{ when } x_i = \arg\max_{x_i \in s_i} \rho(x_i, \ \hat{\theta}(x_1,\ldots,x_n))$ 

- the minimax logistic regression with one interval-censored covariate can always be obtained by computing at most two conventional logistic regressions (for the two extreme choices of compatible precise data sets)
- in the logistic regression with the (unknown) precise covariate, the increase over time of the debris ingestion probability is statistically significant (p-value < 0.001) according to the LR test, because the same is true in the conventional (minimax) logistic regression with worst-case precise data values



 $\bullet \ \hat{\theta}_{\min\max}(s_1,\ldots,s_n) = \hat{\theta}(x_1,\ldots,x_n) \text{ when } x_i = \arg\max_{x_i \in s_i} \rho(x_i, \ \hat{\theta}(x_1,\ldots,x_n))$ 

- the minimax logistic regression with one interval-censored covariate can always be obtained by computing at most two conventional logistic regressions (for the two extreme choices of compatible precise data sets)
- in the logistic regression with the (unknown) precise covariate, the increase over time of the debris ingestion probability is statistically significant (p-value < 0.001) according to the LR test, because the same is true in the conventional (minimax) logistic regression with worst-case precise data values
- note however that this reasoning is not valid in general for other tests (e.g., the Wald test)

M-estimation can be adapted to the case of incomplete/censored data

- M-estimation can be adapted to the case of incomplete/censored data
- ▶ in particular, the minimax M-estimation:

- M-estimation can be adapted to the case of incomplete/censored data
- ▶ in particular, the minimax M-estimation:
  - generalises the nonparametric/pragmatic approach of finding the best fit to the available data, without assuming the existence of a true model

- M-estimation can be adapted to the case of incomplete/censored data
- ▶ in particular, the minimax M-estimation:
  - generalises the nonparametric/pragmatic approach of finding the best fit to the available data, without assuming the existence of a true model
  - can often be implemented as a conventional M-estimation with worst-case precise data values, allowing also statistical inferences about the distribution of the true (but not completely known) precise data values

- M-estimation can be adapted to the case of incomplete/censored data
- ▶ in particular, the minimax M-estimation:
  - generalises the nonparametric/pragmatic approach of finding the best fit to the available data, without assuming the existence of a true model
  - can often be implemented as a conventional M-estimation with worst-case precise data values, allowing also statistical inferences about the distribution of the true (but not completely known) precise data values
  - can be slightly generalised to include also, e.g., minimax Least Quantile of Squares regression (Cattaneo and Wiencierz, 2012, 2014), or minimax Support Vector Regression (Utkin and Coolen, 2011; Wiencierz and Cattaneo, 2015)

- M-estimation can be adapted to the case of incomplete/censored data
- ▶ in particular, the minimax M-estimation:
  - generalises the nonparametric/pragmatic approach of finding the best fit to the available data, without assuming the existence of a true model
  - can often be implemented as a conventional M-estimation with worst-case precise data values, allowing also statistical inferences about the distribution of the true (but not completely known) precise data values
  - can be slightly generalised to include also, e.g., minimax Least Quantile of Squares regression (Cattaneo and Wiencierz, 2012, 2014), or minimax Support Vector Regression (Utkin and Coolen, 2011; Wiencierz and Cattaneo, 2015)
- by contrast, the parametric/idealistic approach of finding the hypothetical true model can be pursued by minimin M-estimation, but is often severely hindered by partial identification

#### references

- Cattaneo, M., and Wiencierz, A. (2012). Likelihood-based Imprecise Regression. Int. J. Approx. Reasoning 53, 1137–1154.
- Cattaneo, M., and Wiencierz, A. (2014). On the implementation of LIR: the case of simple linear regression with interval data. *Comput. Stat.* 29, 743–767.
- Huber, P. J., and Ronchetti, E. M. (2009). Robust Statistics. 2nd edn. Wiley.

Manski, C. F. (2003). Partial Identification of Probability Distributions. Springer.

- Schuyler, Q., Hardesty, B. D., Wilcox, C., and Townsend, K. (2014). Global analysis of anthropogenic debris ingestion by sea turtles. *Conserv. Biol.* 28, 129–139.
- Utkin, L. V., and Coolen, F. P. A. (2011). Interval-valued regression and classification models in the framework of machine learning. In *ISIPTA '11*, eds.
  F. Coolen, G. de Cooman, T. Fetz, and M. Oberguggenberger. SIPTA, 371–380.
- Wiencierz, A., and Cattaneo, M. (2015). On the validity of minimin and minimax methods for Support Vector Regression with interval data. In *ISIPTA '15*, eds. T. Augustin, S. Doria, E. Miranda, and E. Quaeghebeur. Aracne, 325–332.